

Random algebraic polynomials with non-identically distributed normal coefficients

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Abstract

The asymptotic estimate for the expected number of real zeros of a random

$$\text{algebraic polynomial } Q_n(x, \omega) = \sum_{j=0}^n a_j(\omega) \binom{n}{j}^{1/2} x^j,$$

Where the coefficients $a_j(\omega)$ are independently normally distributed random variables defined on a probability space (Ω, F, P) , $\omega \in \Omega$ is known. In this paper we provide the asymptotic value for the expected number of zeros of the integration of $Q_n(x, \omega)$ with respect to x .

Key words & Phrases: Random algebraic polynomial, Expected number of zeroes

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1. Introduction

Let us consider the random algebraic polynomial as

$$Q_n(x, \omega) = \sum_{j=0}^n a_j(\omega) \binom{n}{j}^{1/2} x^j \quad (1.1)$$

Where $a_j(\omega)$, $j=0,1,2,\dots,n$, $\omega \in \Omega$,

is a sequence of independent random variables defined on a probability space (Ω, F, P) . Then

$$F_n(t, \omega) = F_n(t) = \int_0^t Q_n(x, \omega) dx \quad (1.2)$$

$$= \sum_{j=0}^n a_j(\omega) \binom{n}{j}^{1/2} \frac{t^{j+1}}{j+1}$$

represents the area under the curve represen-

ting $Q_n(x, \omega)$ between 0 and $|t|$. Considering the coefficients to be distributed as normal with mean zero and variance unity, Edelman and Kostlan² were first to show that the expected number of real zeros of $Q_n(x, \omega)$ in $(-\infty, \infty)$ is asymptotic to $\sqrt{n}$. Later on, the expected number of K -level crossings, the number of maxima and minima or the number of maxima (minima) below (above) a level are studied by Farahmand in³ and ⁴. Further the number of zeros for the area under curve representing a polynomial of the form as in¹ is studied by Farahmand and Jahangiri⁵ and showed that it is also equal to $\sqrt{n}$. Now we are motivated to find out the number of zeros for the area under the curve representing the same polynomial, where the coefficients are independently normally distributed random

variables with non zero mean, $E(a_j) = \mu \binom{n}{j}^{1/2}$, $\mu \neq 0$ and variance, $V(a_j) = \sigma^2$. The particular choice of polynomial in (1.1) is of special interest for physical properties as described by Ramponi. The analysis employed here may be utilized to find out expected number of zeros for other type of non normal distributions. Hence we have the following theorem⁶.

Theorem : If the coefficients a_j of $F_n(t, \omega)$ in (2) are independently normally distributed random variables with expected

value $E(a_j) = \mu \binom{n}{j}^{1/2}$, $\mu \neq 0$, and variance,

$V(a_j) = \sigma^2$, then the expected number of real zeros of $F_n(t, \omega)$ satisfies

$$EN_F(-\infty, \infty) = \left(\frac{n+1}{2\pi} \right) \log \sqrt{2n}$$

2. Moments:

First of all we need to obtain the estimates for the mean of $F_n(t)$, the mean of its derivative with respect to t , $F_n'(t)$, the variance of $F_n(t)$, the variance of its derivative with respect to t , $F_n'(t)$, and the covariance of $F(t)$ and $F_n'(t)$. From the definition of $F_n(t)$ in (1.2) we have.

$$\begin{aligned} m = E(F_n(t)) &= \mu \sum_{j=0}^n \binom{n}{j} \frac{t^{j+1}}{j+1} \quad (2.1) \\ &= \mu \frac{(1+t)^{n+1}}{n+1} \end{aligned}$$

$$\begin{aligned} m' = E(F_n'(t)) &= \mu \sum_{j=0}^n \binom{n}{j} t^j \quad (2.2) \\ &= \mu (1+t)^n \end{aligned}$$

$$\begin{aligned} A^2 = V(F_n(t)) &= \sigma^2 \sum_{j=0}^n \binom{n}{j} \frac{t^{2j+2}}{(j+1)^2} \quad (2.3) \\ &\sim \sigma^2 \frac{(1+t^2)^{n+2}}{(n+1)^2} \left[1 + O\left(\frac{1}{n}\right) \right] \end{aligned}$$

$$\begin{aligned} B^2 = V(F_n'(t)) &= \sigma^2 \sum_{j=0}^n \binom{n}{j} t^{2j} \quad (2.4) \\ &= \sigma^2 (1+t^2)^n \end{aligned}$$

$$\text{(Using the binomial identity } \sum_{j=0}^n \binom{n}{j} t^j = (1+t)^n \text{)}$$

$$C = \text{Cov}(F_n(t), F_n'(t)) = \sigma^2 \sum_{j=0}^n \binom{n}{j} \frac{t^{2j+1}}{(j+1)} \quad (2.5)$$

$$= \sigma^2 \frac{(1+t^2)^{n+1}}{(n+1)t}$$

(By integrating the above binomial identity)
Then from Cramer and Leadbetter, under the condition that $A^2 B^2 > C^2$, the expected number of real zeros of $F_n(t)$ in the interval (α, β) as,

$$EN_F(\alpha, \beta) = \int_{\alpha}^{\beta} \frac{\Delta}{A^2} \varphi\left(\frac{m}{\sigma^2}\right) [2\varphi(\eta) + \eta(2\Phi(\eta) - 1)] dt \quad (2.6)$$

Where

$$\Delta^2 = A^2 B^2 - C^2$$

$$\eta = \left(\frac{A^2 m' - Cm}{A\Delta} \right)$$

$$\varphi(t) = \frac{1}{\sqrt{2\pi}} e^{-\left(\frac{t^2}{2}\right)}$$

$$\Phi(t) = \frac{1}{\sqrt{2\pi}} \int_{-\infty}^t \varphi(y) dy$$

Assuming $\text{erf}(x) = \int_0^x e^{-t^2} dt$, the equation (2.6) can be written as,

$$EN_F(\alpha, \beta) = I_1(\alpha, \beta) + I_2(\alpha, \beta)$$

Where,

$$I_1(\alpha, \beta) = \int_{\alpha}^{\beta} \frac{\Delta}{\pi A^2} e^{\left[-\left(\frac{m^2}{2A^2}\right) + \left(\frac{A^2 m' - Cm}{\sqrt{2}A\Delta}\right)^2 \right]} dt \quad (2.7)$$

$$I_2(\alpha, \beta) = \int_{\alpha}^{\beta} \left[\frac{\sqrt{2}(A^2 m' - Cm)}{\pi A^3} e^{\left(\frac{m^2}{2A^2}\right)} \text{erf}\left(\frac{A^2 m' - Cm}{\sqrt{2}A\Delta}\right) \right] dt \quad (2.8)$$

3. Proof of the theorem :

Now to estimate $I_2(-\infty, \infty)$, let us assume a new variable 'u' as ,

$$u = \frac{m}{A}$$

$$= \frac{\mu t(1+t)^{n+1}}{\sigma(1+t^2)^{\frac{n+2}{2}} \left[1 + O\left(\frac{1}{n}\right) \right]} \quad (3.1)$$

$$\approx \frac{\mu t(1+t)^{n+1}}{\sigma(1+t^2)^{\frac{n+2}{2}}}$$

for sufficiently large n, $\left[1 + O\left(\frac{1}{n}\right) \right] \rightarrow 1$

Hence,

$$\frac{du}{dt} = \frac{\mu}{\sigma} (1+t^2)^{-\left(\frac{n+4}{2}\right)} (1+t)^n (1+nt+2t-t^2-nt^2) \quad (3.2)$$

On Simplification,

$$\left(\frac{A^2 m' - Cm}{A^3} \right) = \frac{\mu t (1+t)^n (1+t^2)^{-\left(\frac{n+4}{2}\right)} (1-t)}{\sigma} \quad (3.3)$$

We have , from (2.8) , (3.1) , (3.2) and (3.3)

$$\begin{aligned} I_2(-\infty, \infty) &\leq \int_{-\infty}^{\infty} \left(\frac{A^2 m' - Cm}{\sqrt{2\pi} \Delta^3} \right) e^{\left(-\frac{m^2}{2A^2} \right)} dt \\ &\leq \frac{1}{\sqrt{2\pi}} \int_{-\infty}^{\infty} e^{\left(-\frac{u^2}{2} \right)} du \times O(1) \end{aligned}$$

$$= O(1) \quad (3.4)$$

Now to estimate $I_1(-\infty, \infty)$, let us assume,

$$\begin{aligned} \tau(t) &= \frac{m^2}{2A^2} + \left(\frac{A^2 m' - Cm}{\sqrt{2} A \Delta} \right)^2 \\ &= \frac{\frac{\mu^2 (1+t)^{2n+2}}{(n+1)^2}}{2\sigma^2 (1+t^2)^{n+2} \left[1 + O\left(\frac{1}{n} \right) \right]} \\ &= \frac{\mu^2 (1+t)^{2n+2}}{t^2 (n+1)^2} \end{aligned}$$

$$+ \left[\frac{\sigma^2 (1+t^2)^{n+2} \left[1 + O\left(\frac{1}{n} \right) \right] \mu (1+t)^n}{t^2 (n+1)^2} - \frac{\sigma^2 (1+t^2)^{n+1} \mu (1+t)^{n+1} (1+t)^{2n} (2+t^2-t)}{t(n+1) (n+1)} \right. \\ \left. + \frac{\sqrt{2} \sigma^2 (1+t^2)^{\frac{n+2}{2}} \left[1 + O\left(\frac{1}{n} \right) \right] \sigma^2 (1+t^2)^{n+1} O\left(\frac{1}{\sqrt{n}} \right)}{t(n+1)} \right]$$

For sufficiently large n, $O\left(\frac{1}{\sqrt{n}} \right) \rightarrow 0$, So

second term vanishes

$$\begin{aligned} \text{Now } \Delta^2 &= A^2 B^2 - C^2 \\ &= \frac{\sigma^4 (1+t^2)^{2n+2}}{t^2 (n+1)^2} O\left(\frac{1}{n} \right) \end{aligned}$$

$$\tau(t) = \frac{\mu^2 t^2 (1+t)^{2n+2}}{2\sigma^2 (1+t^2)^{n+2}}$$

So,

$$= \frac{\mu^2}{2\sigma^2} \left(\frac{t^2}{1+t^2} \right) \left(\frac{(1+t)^2}{(1+t^2)} \right)^{(n+1)} \quad (3.5)$$

$$\Delta = \frac{\sigma^2 (1+t^2)^{n+1}}{t (n+1)} O\left(\frac{1}{\sqrt{n}} \right) \quad (3.6)$$

So that,

$$\frac{\Delta}{A^2} = \frac{\sqrt{nt}}{(1+t^2)} \quad (3.7)$$

Now Consider $0 < t < 1$, the integral $I_1(0,1)$ can be written as

$$I_1(0, 1) = I_1(0, \varepsilon) + I_1(\varepsilon, 1) \quad (3.8)$$

Where $\varepsilon = \frac{1}{\sqrt{n}} \rightarrow 0$, as $n \rightarrow \infty$

Then,

$$\begin{aligned} I_1(0, \varepsilon) &= \int_0^\varepsilon \frac{\Delta}{\pi A^2} \exp(-\tau(t)) dt \\ &\leq \frac{\sqrt{n}}{\pi} \int_0^\varepsilon \frac{t}{(1+t^2)} dt \quad (\text{Where } \tau(t) \rightarrow 0, \text{ as } n \rightarrow \infty) \\ &= \frac{\sqrt{n}}{\pi} [\log(1+t^2)]_0^\varepsilon \\ &= \frac{\sqrt{n}}{\pi} \log\left(1 + \frac{1}{n}\right) \quad (\text{Choosing } \varepsilon = \frac{1}{\sqrt{n}}) \\ &\rightarrow 0, \text{ as } n \rightarrow \infty \end{aligned}$$

So,

$$I_1(0, \varepsilon) \rightarrow 0 \text{ as } n \rightarrow \infty \quad (3.9)$$

Next consider $\varepsilon < t < 1$,

$$\text{So, } I_1(\varepsilon, 1) = \int_\varepsilon^1 \frac{\Delta}{\pi A^2} e^{(-\tau(t))} dt$$

$$= \frac{\sqrt{n}}{\pi} \int_\varepsilon^1 \frac{t}{(1+t^2)} e^{(-\tau(t))} dt$$

$\rightarrow 0$, as $\tau(t) \rightarrow \infty$, where $n \rightarrow \infty$

Therefore ,

$$I_1(\varepsilon, 1) \rightarrow 0 \text{ as } n \rightarrow \infty \quad (3.10)$$

Hence from (3.9) and (3.10), we obtain

$$I_1(0,1) \rightarrow 0 \text{ as } n \rightarrow \infty \quad (3.11)$$

Then consider $-1 < t < 0$, the integral, $I_1(-1, 0)$ can be written as

$$I_1(-1, 0) = I_1(-1, -\varepsilon) + I_1(-\varepsilon, 0) \quad (3.12)$$

So

$$\begin{aligned} I_1(-\varepsilon, 0) &= \int_{-\varepsilon}^0 \frac{\Delta}{\pi A^2} e^{(-\tau(t))} dt \\ &\leq \frac{\sqrt{n}}{\pi} \int_{-\varepsilon}^0 \frac{t}{(1+t^2)} dt \quad (\text{Where } \tau(t) \rightarrow 0, \text{ as } n \rightarrow \infty) \\ &= \frac{\sqrt{n}}{\pi} [\log(1+t^2)]_{-\varepsilon}^0 \\ &= \frac{-\sqrt{n}}{\pi} \log(1+\varepsilon^2) \\ &= \frac{-\sqrt{n}}{\pi} \log\left(1 + \frac{1}{n}\right) \quad (\text{Choosing } \varepsilon = \frac{1}{\sqrt{n}}) \\ &\rightarrow 0 \text{ as } n \rightarrow \infty \end{aligned}$$

So,

$$I_1(-\varepsilon, 0) \rightarrow 0 \text{ as } n \rightarrow \infty \quad (3.13)$$

Next consider $-1 < t < -\varepsilon$, We have,

$$\tau(t) = \frac{\mu^2}{4\sigma^2} \left(\frac{t^2}{1+t^2} \right) \left(\frac{(1+t)^2}{(1+t^2)} \right)^{(n+1)}$$

Putting $t=-y$, so that $\varepsilon < y < 1$, then $\tau(t)$ becomes $\tau_1(y)$ as

$$\tau_1(y) = \frac{\mu^2}{4\sigma^2} \left(\frac{y^2}{1+y^2} \right) \left(\frac{(1-y)^2}{(1+y^2)} \right)^{(n+1)} \quad (3.14)$$

Where

$$\tau_1(y) \rightarrow 0 \text{ as } n \rightarrow \infty$$

So for

$$-1 < t < -\varepsilon$$

$$\begin{aligned} I_1(-1, -\varepsilon) &= \int_{-1}^{-\varepsilon} \frac{\Delta}{\pi A^2} e^{(-\tau(t))} dt \\ &= \frac{\sqrt{n}}{\pi} \int_{-1}^{-\varepsilon} \frac{t}{(1+t^2)} e^{(-\tau(t))} dt \end{aligned}$$

For $\varepsilon < y < 1$

$$\text{Now } I_1(\varepsilon, 1) = \frac{\sqrt{n}}{\pi} \int_{\varepsilon}^1 \frac{y}{(1+y^2)} e^{(-\tau_1(y))} dy$$

$$= \frac{\sqrt{n}}{\pi} \int_{\varepsilon}^1 \frac{y}{(1+y^2)} dy \quad (\text{For sufficiently large } n)$$

$$\begin{aligned} &= \frac{\sqrt{n}}{\pi} \left[\log(1+y^2) \right]_{\varepsilon}^1 \\ &= \frac{\sqrt{n}}{\pi} \log \left(\frac{2n}{n+1} \right)^{\frac{1}{2}} \end{aligned} \quad (3.15)$$

Then from (3.13) & (3.15) it follows that

$$I_1(-1, 0) = \frac{\sqrt{n}}{\pi} \log \left(\frac{2n}{n+1} \right)^{\frac{1}{2}} \quad (3.16)$$

Consider $1 < t < \infty$, put $t = \frac{1}{y}$, So that $0 < y < 1$

$$\begin{aligned} \text{For } 1 < t < \infty, \quad I_1(1, \infty) &= I_1(0, 1), \text{ For } 0 < y < 1 \\ &= I_1(0, \varepsilon) + I_1(\varepsilon, 1) \end{aligned}$$

We have

$$\tau(t) = \frac{\mu^2}{4\sigma^2} \left(\frac{t^2}{1+t^2} \right) \left(\frac{(1+t)^2}{(1+t^2)} \right)^{(n+1)}$$

For $t = \frac{1}{y}$, $\tau(t)$ becomes $\tau_2(y)$ as

$$\tau_2(y) = \frac{\mu^2}{4\sigma^2} \frac{1}{(1+y^2)} \left(\frac{(1+y)^2}{(1+y^2)} \right)^{(n+1)} \quad (3.17)$$

For $1 < t < \infty$,

$$I_1(1, \infty) = \int_1^{\infty} \frac{\Delta}{\pi A^2} e^{(-\tau(t))} dt$$

$$= \frac{\sqrt{n}}{\pi} \int_1^{\infty} \frac{t}{(1+t^2)} e^{(-\tau(t))} dt \quad \text{For } -1 < y < 0$$

For $0 < y < 1$

Now $I_1(0,1) = I_1(0, \varepsilon) + I_1(\varepsilon, 1)$

$$\begin{aligned} I_1(0, \varepsilon) &= \frac{-(n+1)}{2\pi} \int_0^{\varepsilon} \frac{1}{y(1+y^2)} e^{\left(-\tau_2(y)\right)} dy \\ &= \frac{(n+1)}{2\pi} \log \sqrt{n+1} \end{aligned} \quad (3.18)$$

$$\begin{aligned} I_1(\varepsilon, 1) &= \frac{-(n+1)}{2\pi} \int_{\varepsilon}^1 \frac{1}{y(1+y^2)} e^{\left(-\tau_2(y)\right)} dy \\ &\rightarrow 0 \text{ as } \tau_2(y) \rightarrow \infty \text{ as For sufficiently large } n \end{aligned} \quad (3.19)$$

Then from (3.18) and (3.19), it follows that

$$I_1(1, \infty) = \frac{\sqrt{n}}{\pi} \log \sqrt{n+1} \quad (3.20)$$

Consider $-\infty < t < -1$, put $y = \frac{1}{t}$, So that $-1 < y < 0$,

The Integral $I_1(-\infty, -1)$ for $-\infty < t < -1$ can be written as $I_1(-1, 0)$ for $-1 < y < 0$

For $-\infty < t < -1$

$$I_1(-\infty, -1) = \frac{\sqrt{n}}{\pi} \int_{-\infty}^{-1} \frac{t}{(1+t^2)} e^{(-\tau(t))} dt$$

$$I_1(-1, 0) = \frac{-\sqrt{n}}{\pi} \int_{-1}^0 \frac{1}{y(1+y^2)} e^{(-\tau_2(y))} dy$$

Using (3.17)

$$= \frac{-\sqrt{n}}{\pi} \int_{-1}^0 \frac{1}{y(1+y^2)} dy = dy \text{ (as } \tau_2(y) \rightarrow 0 \text{ as } n \rightarrow \infty)$$

$$= \frac{-\sqrt{n}}{\pi} \int_{\frac{3\pi}{4}}^0 \frac{1}{\tan \theta} d\theta$$

(Putting $y = \tan \theta \Rightarrow dy = \sec^2 \theta d\theta$)

$$= \frac{-\sqrt{n}}{\pi} [\log \sin \theta]_{\frac{3\pi}{4}}^0$$

$$= \frac{-\sqrt{n}}{\pi} \log \frac{1}{\sqrt{2}}$$

$$= 0 \text{ (Assuming } \log 0 = 0) \quad (3.21)$$

Therefore from (3.11), (3.16), (3.20) and (3.21) it follows that

$$I_1(-\infty, \infty) = \frac{-\sqrt{n}}{\pi} \log \left(\frac{2n}{n+1} \right)^{\frac{1}{2}}$$

Hence

$$EN_F(-\infty, \infty) = I_1(-\infty, \infty) + I_2(-\infty, \infty)$$

$$\sim \frac{\sqrt{n}}{\pi} \log \left(\frac{2n}{n+1} \right)^{\frac{1}{2}}$$

Which proves the theorem.

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