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The Batch Service Vacation Model $(M/G^{(a,b)}/1): (E, SV)$:

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Abstract

In this chapter we discuss the batch service vacation model. In such system, customers arrive according to a Poisson process and are served in batches of maximum size b and minimum threshold a . The server takes a single vacation when it finds less than a customers after the service completion.

Key words: Batch Service, Single Vacation, Poisson Process.

Introduction

Consider the $(M/G^{(a,b)}/1)$ queue with batch

service vacation model. The steady state of the system can be described by the following random variable.

$$\xi = \begin{cases} 0 & \text{If the server is dormant and ready to serves} \\ 1 & \text{If the server is on vacation,} \\ 2 & \text{If the server is busy} \end{cases}$$

L_s = The number of customers Present in the system,

T = The residual service time of the batch in service,

V = The residual vacation time for the server on vacation.

There are difference in the definitions of ξ and T between the Batch service Model and the batch arrival Model. We define

$$\pi_n(x) dx = P(L_s = n, x < T \leq x + dx, \xi = 2), \quad n = 0, 1, 2, \dots$$

$$w_n(x) dx = P(L_s = n, x < V \leq x + dx, \xi = 1), \quad n = 0, 1, 2, \dots$$

$$R_n = P(L_s = n, \xi = 0) \quad n = 0, 1, 2, \dots, a-1$$

And the LST

$$\pi_n^*(s) = \int_0^\infty e^{-sx} \pi_n(x) dx,$$

$$w_n^*(s) = \int_0^\infty e^{-sx} w_n(x) dx$$

It follows from the above that

$$\pi_n^*(0) = \pi_n = \int_0^\infty \pi_n(x) dx \quad \text{and} \quad w_n^*(0) = w_n = \int_0^\infty w_n(x) dx$$

It is clear that $\pi_n(w_n), n \geq 0$, represent the Prob. of n customers in the queue when the server is busy at arbitrary time instants.

By considering the steady-state transitions, we obtain the following system of differential difference equations.

$$0 = -\lambda R_0 + W_0(0)$$

$$0 = -\lambda R_n + \lambda R_{n-1} + w_n(0), 1 \leq n \leq a-1 \quad (A)$$

$$-\frac{d\pi_0(x)}{dx} = -\lambda \pi_0(x) + b(x) \sum_{n=a}^b [\pi_n(0) + w_n(0)] + \lambda R_{a-1} b(x), \quad (1)$$

$$-\frac{d\pi_n(x)}{dx} = -\lambda\pi_n(x) + \lambda_{n-1}(x) + b(x) [\pi_{n+b}(o) + w_{n+b}(o)] \quad n \geq 1 \quad (2)$$

$$-\frac{dw_0(x)}{dx} = -\lambda w_0(x) + \pi_0(o) v(x), \quad (3)$$

$$-\frac{dw_n(x)}{dx} = -\lambda w_n(x) + \lambda w_{n-1}(x) + \pi_n(o) v(x), \quad 1 \leq n \leq a-1, \quad (4)$$

$$-\frac{dw_n(x)}{dx} = -\lambda w_n(x) + \lambda w_{n-1}(x), \quad n \geq a \quad (5)$$

Taking the LST on both sides of the above equation (1) to (5).

$$(\lambda - S)\pi_0^*(s) = T^*(s) \sum_{n=a}^b [\pi_n(o) + w_n(o)] + \lambda R_{a-1} T^*(s) - \pi_0(o) \quad (6)$$

$$(\lambda - S)\pi_n^*(s) = \lambda \pi_{n-1}^*(s) + T^*(s) [\pi_{n+b}(o) + w_{n+b}(o)] - \pi_n(o), \quad n \geq 1 \quad (7)$$

$$(\lambda - S)w_0^*(s) = v^*(s) \pi_0(o) - w_0(o) \quad (8)$$

$$(\lambda - S)w_n^*(s) = \lambda w_{n-1}^*(s) + V^*(s) \pi_n(o) - w_n(o), \quad 1 \leq n \leq a-1 \quad (9)$$

$$(\lambda - S)w_n^*(s) = \lambda w_{n-1}^*(s) - w_n(o), \quad n \geq a \quad (10)$$

Now using equation (1) & (2) and equation (6) to (10) we obtain a set of Results that the later lead to queue length distribution at various epochs.

Lemma 1:- There exist two relations

$$\sum_{n=0}^j w_n(o) = \lambda R_j \quad 0 \leq j \leq a-1 \quad (11)$$

$$\sum_{n=0}^{a-1} \pi_n(o) = \sum_{n=0}^{\infty} w_n(o) \quad (12)$$

Proof: From equation (A) and letting $n=1$, we obtain

$$\sum_{n=0}^1 w_n(o) = \lambda R_1 \quad n = 0, 1, 2, \dots, a-1$$

We get equation (11) of lemma 1.

Setting $S=0$ from equation (6) & (7) we get

$$\pi_0(o) = \sum_{n=a}^b [\pi_n(o) + w_n(o)] + \lambda R_{a-1} - \lambda \pi_0 \quad (13)$$

$$\pi_0(o) = \pi_{n+b}(o) + w_{n+b}(o) + \lambda (\pi_{n-1} - \pi_n) \quad n \geq 1 \quad (14)$$

Summing over n equation (14) adding equation (13) and using equation (11) we obtain equation (12).

Define the non-serving period D_V^c as the sum of vacation V and on Idle time I_V , we have the following Lemma.

Lemma 2: The Expected value of D_V^c is given by

$$E(D_V^c) = E(v) + \frac{1}{\sum_{n=0}^{a-1} P_n^+} \left[\sum_{i=0}^{a-1} P_i^+ \sum_{j=0}^{a-1-i} h_j \frac{(a-i-j)}{\lambda} \right] \quad (15)$$

Where P_i^+ is the stationary probability that i customers are left at a departure instant of a batch, and

$$h_j = \int_0^{\infty} \frac{(\lambda x)^j}{j!} e^{-\lambda x} dv(x)$$

Proof: Let $N(t)$ [The number of customers in the system at time (t)] be the state of system at time t . Thus, at the end of busy period, $N(t)$ enters the set of vacation states $S = \{0, 1, 2, \dots, a-1\}$. The conditional probability that $N(t)$ enters state $i \in S$, given that $N(t)$ enters, then

$$S = P_i^+ / \sum_{n=0}^{a-1} P_n^+$$

For fixed $i \in S$, if $j \leq (a-1-i)$ customers arrive during a vacation with probability h_j , then at the vacation completion instant, $N(t)$ enters the set of idle states $U = \{K : k = a-i-j\}$. That $N(t)$ leaves the Set U when $a-(i+j)$ customers arrive. Thus the expected time of $N(t)$ is in

$$(a-i-j)/\lambda$$

Using the conditional Argument, Then

$$E(D_V^c) = E(V) + E(I_V)$$

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